

Progress in the method of Ghosts and Shadows for Beta Ensembles



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Wishart Matrices (arbitrary covariance)

- $G = m \times n$ matrix of Gaussians
- $\Sigma = m \times n$ semidefinite matrix
- $G'G \Sigma$ is similar to $A = \Sigma^{1/2} G'G \Sigma^{-1/2}$
- For $\beta = 1, 2, 4$, the joint eigenvalue density of A has a formula:

$$C_{m,n,\Sigma,\beta} \prod_{i=1}^n \lambda_i^{\frac{m-n+1}{2}\beta-1} \prod_{j < k} |\lambda_k - \lambda_j|^\beta \cdot {}_0F_0^{(\beta)}\left(-\frac{\beta}{2}\Lambda, \Sigma^{-1}\right),$$

where $\Lambda = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_n)$.

- Known for $\beta = 2$ in some circles as Harish-Chandra-Itzykson-Zuber

Main Purpose of this talk

$$C_{m,n,\Sigma,\beta} \prod_{i=1}^n \lambda_i^{\frac{m-n+1}{2}\beta-1} \prod_{j<k} |\lambda_k - \lambda_j|^\beta \cdot {}_0F_0^{(\beta)}\left(-\frac{\beta}{2}\Lambda, \Sigma^{-1}\right),$$

where $\Lambda = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_n)$.

- Eigenvalue density of $G'G \Sigma$ (similar to $A=\Sigma^{1/2}G'G\Sigma^{-1/2}$)
- Present an algorithm for sampling from this density
- Show how the method of Ghosts and Shadows can be used to derive this algorithm
- Further evidence that $\beta=1,2,4$ need not be special

Eigenvalues of GOE ($\beta=1$)

- Naïve Way:

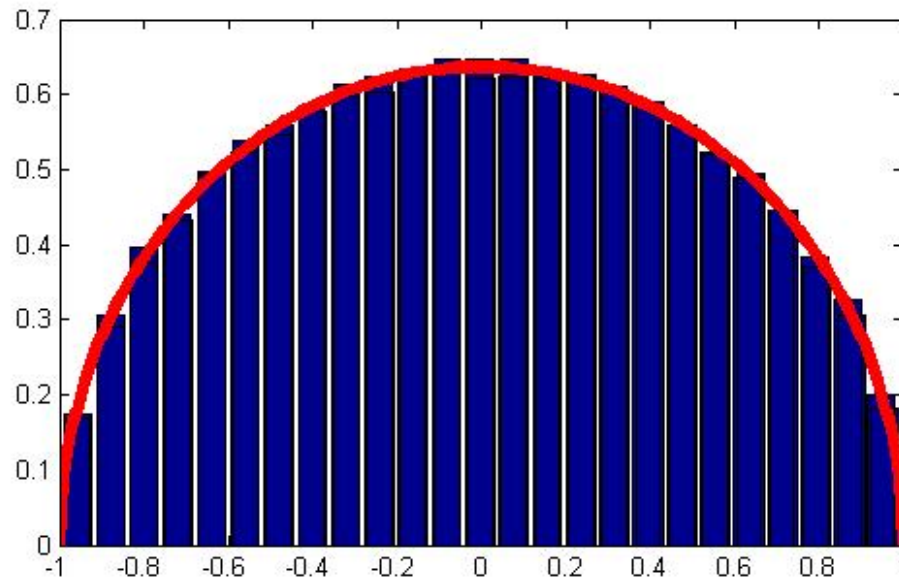
MATLAB[®]: `A=randn(n); S=(A+A')/sqrt(2*n); eig(S)`

R:

`A=matrix(rnorm(n*n),ncol=n); S=(a+t(a))/sqrt(2*n); eigen(S,symmetric=T,only.values=T)$values;`

Mathematica:

`A=RandomArray[NormalDistribution[],{n,n}]; S=(A+Transpose[A])/Sqrt[n]; Eigenvalues[s]`



Introduction to Ghosts

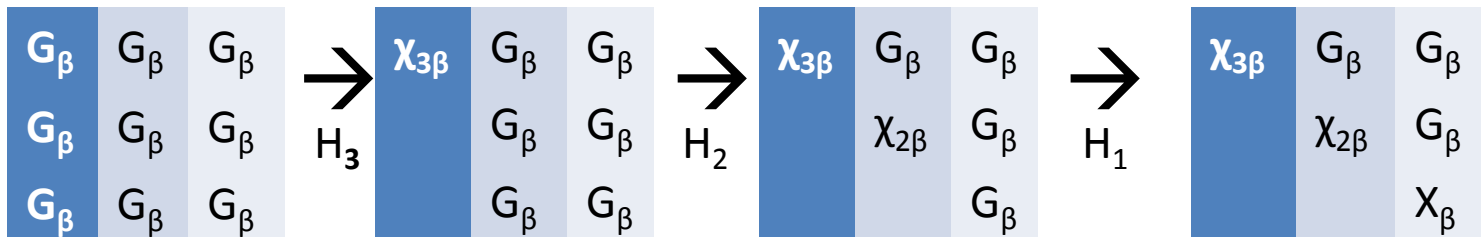
- G_1 is a standard normal $N(0,1)$
- G_2 is a complex normal $(G_1 + iG_1)$
- G_4 is a quaternion normal $(G_1 + iG_1 + jG_1 + kG_1)$
- G_β ($\beta > 0$) seems to often work just fine
→ “Ghost Gaussian”

Chi-squared

- Defn: χ_β^2 is the sum of β iid squares of standard normals if $\beta=1,2,\dots$
- Generalizes for non-integer β as the “gamma” function interpolates factorial
- χ_β is the sqrt of the sum of squares (which generalizes) (wikipedia chi-distribution)
- $|G_1|$ is χ_1 , $|G_2|$ is χ_2 , $|G_4|$ is χ_4
- So why not $|G_\beta|$ is χ_β ?
- I call χ_β the shadow of G_β

Linear Algebra Example

- Given a ghost Gaussian, G_β , $|G_\beta|$ is a real chi-beta variable χ_β variable.
- Gram-Schmidt:

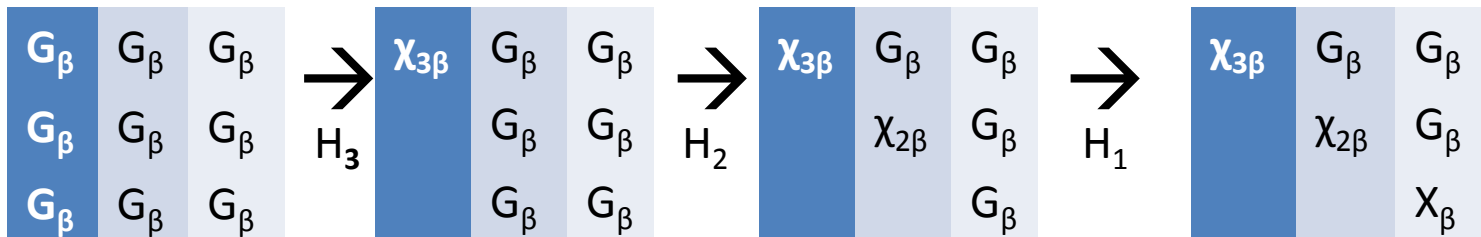


or

$$\begin{bmatrix} G_\beta & G_\beta & G_\beta \\ G_\beta & G_\beta & G_\beta \\ G_\beta & G_\beta & G_\beta \end{bmatrix} = [Q] * \begin{bmatrix} \chi_{3\beta} & G_\beta & G_\beta \\ \chi_{2\beta} & G_\beta & G_\beta \\ \chi_\beta & G_\beta & G_\beta \end{bmatrix}$$

Linear Algebra Example

- Given a ghost Gaussian, G_β , $|G_\beta|$ is a real chi-beta variable χ_β variable.
- Gram-Schmidt:



or

$$\begin{bmatrix} G_\beta & G_\beta & G_\beta \\ G_\beta & G_\beta & G_\beta \\ G_\beta & G_\beta & G_\beta \end{bmatrix} = [Q] * \begin{bmatrix} \chi_{3\beta} & G_\beta & G_\beta \\ \chi_{2\beta} & G_\beta & G_\beta \\ \chi_\beta & G_\beta & G_\beta \end{bmatrix}$$

Q has β -Haar Measure!
We have computed moments!

Tridiagonalizing Example

Symmetric Part of

G_β	G_β	G_β	...	G_β
G_β	G_β	G_β	...	G_β
G_β	G_β	G_β	...	G_β
G_β	G_β	G_β	G_β	G_β

$$\rightarrow H_n^\beta = \frac{1}{\sqrt{\beta}} \begin{bmatrix} g_1 & \chi_{(n-1)\beta} & & & \\ \chi_{(n-1)\beta} & g_2 & \chi_{(n-2)\beta} & & \\ & \ddots & \ddots & \ddots & \\ & & \chi_{2\beta} & g_{n-1} & \chi_\beta \\ & & & \chi_\beta & g_n \end{bmatrix}$$

(Silverstein, Trotter, etc)

Histogram without Histogramming: Sturm Sequences

- Count #eigs < 0.5 : Count sign changes in

$$\text{Det}((A-0.5*I)[1:k,1:k])$$

- Count #eigs in $[x,x+h]$

Take difference in number of sign changes at $x+h$ and x

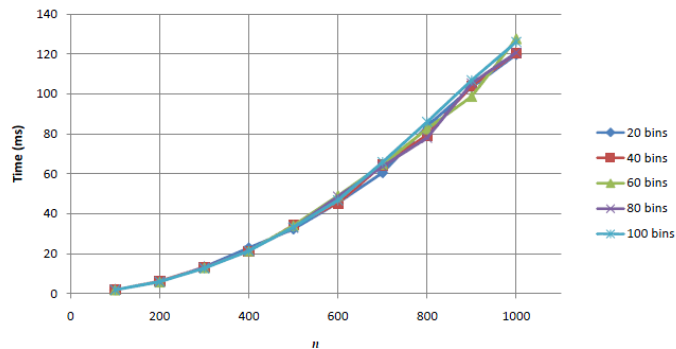


FIGURE 1. Performance of naive histogramming algorithm. This figure makes readily apparent the dominance of the eigenvalue computation (which takes $O(n^2)$ time): the number of histogram bins makes no significant difference in performance.

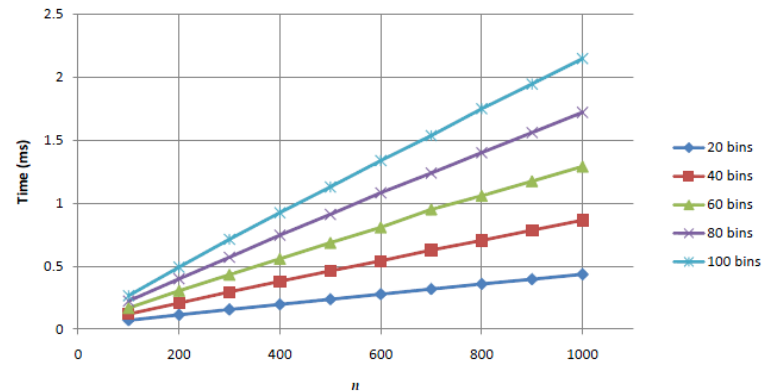


FIGURE 2. Performance of Sturm sequence based histogramming algorithm. This figure clearly demonstrates the bilinear dependency from the $O(mn)$ computation time.

Mentioned in Dumitriu and E 2006, Used theoretically in Albrecht, Chan, and E 2008

A good computational trick is a good theoretical trick!

Corollary 5.3. (“Eigenvalue distribution” or “Level density”) Let $f_{G_{-\lambda}}$ be the Gaussian density with mean $-\lambda$ and variance 1, and let $D_p(z)$ be the parabolic cylinder function. Define $z_i \equiv \text{sign}(s_{i-1})(s_i + \lambda + s_{i-1})$, and $p_i \equiv \frac{1}{2}\beta(i-1)$. If Λ is drawn from the β -Hermite eigenvalue density, then

$$\Pr[\Lambda < \lambda] = \frac{1}{n} \sum_{i=1}^n \int_{-\infty}^0 f_{r_i}(s_i) ds_i,$$

where

$$f_{r_i}(s_i) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} f_{r_1, r_2, \dots, r_n}(s_1, s_2, \dots, s_n) ds_1 ds_2 \dots ds_{i-1},$$

$$f_{r_1, r_2, \dots, r_n}(s_1, s_2, \dots, s_n) = f_{G_{-\lambda}}(s_1) \prod_{i=2}^n f_{r_i|r_{i-1}}(s_i | s_{i-1}), \text{ and}$$

$$f_{r_i|r_{i-1}}(s_i | s_{i-1}) = \frac{|s_{i-1}|^{p_i}}{\sqrt{2\pi}} e^{-\frac{1}{4}[2(s_i + \lambda)^2 - z_i^2]} D_{-p_i}(z_i).$$

Finite Semi-Circle Laws for Any Beta!

Corollary 5.4. (Largest eigenvalue distribution) As in Corollary 5.3, let $f_{G_{-\lambda}}$ be the Gaussian density with mean $-\lambda$ and variance 1, $D_p(z)$ be the parabolic cylinder function, $z_i \equiv \text{sign}(s_{i-1})(s_i + \lambda + s_{i-1})$, and $p_i \equiv \frac{1}{2}\beta(i-1)$. If Λ_{max} is drawn from the β -Hermite largest eigenvalue density, then

$$\Pr[\Lambda_{max} < \lambda] = \int_{-\infty}^0 \int_{-\infty}^0 \cdots \int_{-\infty}^0 f_{r_1, r_2, \dots, r_n}(s_1, s_2, \dots, s_n) ds_1 ds_2 \dots ds_n,$$

where

$$f_{r_1, r_2, \dots, r_n}(s_1, s_2, \dots, s_n) = f_{G_{-\lambda}}(s_1) \prod_{i=2}^n f_{r_i|r_{i-1}}(s_i | s_{i-1}), \text{ and}$$

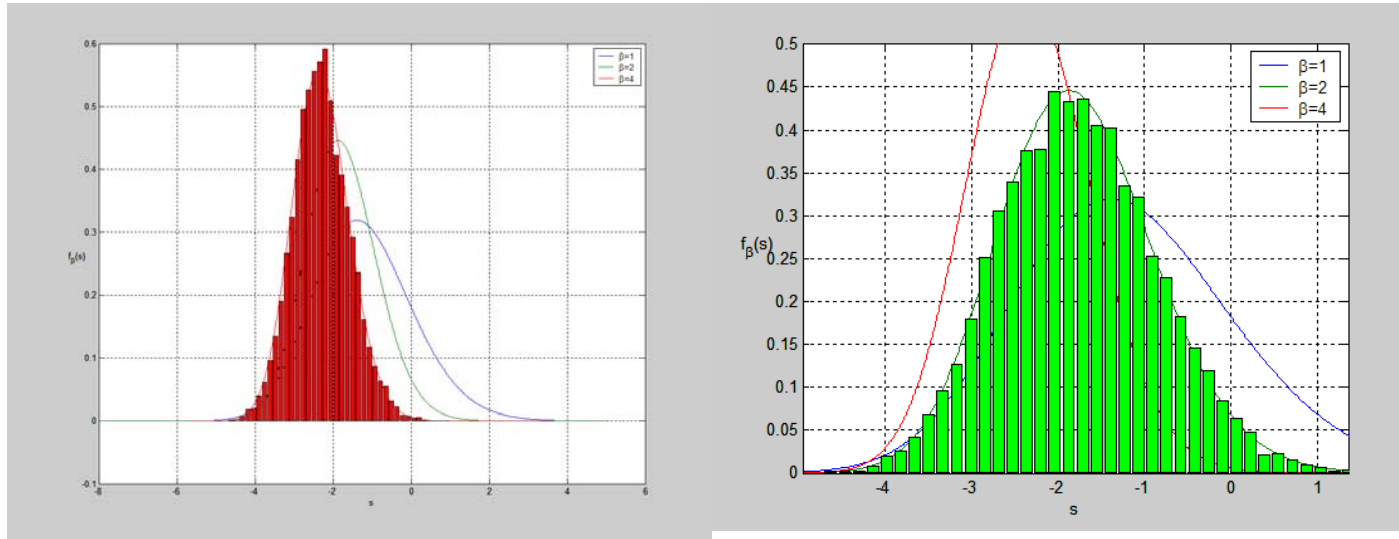
$$f_{r_i|r_{i-1}}(s_i | s_{i-1}) = \frac{|s_{i-1}|^{p_i}}{\sqrt{2\pi}} e^{-\frac{1}{4}[2(s_i + \lambda)^2 - z_i^2]} D_{-p_i}(z_i).$$

Finite Tracy-Widom Laws for Any Beta!

Stochastic Differential Eigen-Equations

- Tridiagonal Models Suggest SDE's with Brownian motion as infinite limit:
 - E: 2002
 - E and Sutton 2005, 2007
 - Brian Rider and (Ramirez, Cambroner, Virag, etc.)
 - (Lots of beautiful results!)
 - (Not today's talk)

“Tracy-Widom” Computations Eigenvalues without the whole matrix!



Never construct the entire tridiagonal matrix!

Just say the upper $10n^{1/3}$ by $10n^{1/3}$

Compute largest eigenvalue of that, perhaps using Lanczos with shift and invert strategy!

Can compute for n amazingly large!!!! And any beta.

E: 2003 Persson and E: 2005

Other Computational Results

- Infinite Random Matrix Theory
 - The Free Probability Calculator (Raj)
- Finite Random Matrix Theory
 - MOPS (Ioana Dumitriu) (β : orthogonal polynomials)
 - Hypergeometrics of Matrix Argument (Plamen Koev) (β : distribution functions for finite stats such as the finite Tracy-Widom laws for Laguerre)
- Good stuff, but not today.

Scary Ideas in Mathematics

- Zero
- Negative
- Radical
- Irrational
- Imaginary
- Ghosts: Something like a commutative algebra of random variables that generalizes random Reals, Complexes, and Quaternions and inspires theoretical results and numerical computation

Did you say “commutative”??

- Quaternions don't commute.
- Yes but random quaternions do!
- If x and y are G_4 then $x*y$ and $y*x$ are identically distributed!

RMT Densities

- Hermite:

$$c \prod |\lambda_i - \lambda_j|^\beta e^{-\sum \lambda_i^2 / 2} \quad (\text{Gaussian Ensemble})$$

- Laguerre:

$$c \prod |\lambda_i - \lambda_j|^\beta \prod \lambda_i^{m_1} e^{-\sum \lambda_i} \quad (\text{Wishart Matrices})$$

- Jacobi:

$$c \prod |\lambda_i - \lambda_j|^\beta \prod \lambda_i^{m_1} \prod (1 - \lambda_i)^{m_2} \quad (\text{Manova Matrices})$$

- Fourier:

$$c \prod |\lambda_i - \lambda_j|^\beta \quad (\text{on the complex unit circle}) \quad \text{“Jack Polynomials”}$$

Traditional Story: Count the real parameters $\beta=1,2,4$ for real, complex, quaternion

Applications: $\beta=1$: All of Multivariate Statistics $\beta=2$: Tons,

$\beta=4$: Almost nobody cares

Dyson 1962 “Threefold Way” Three Associative Division Rings

β -Ghosts

- $\beta=1$ has one real Gaussian (G)
- $\beta=2$ has two real Gaussians ($G+iG$)
- $\beta=4$ has four real Gaussians ($G+iG+jG+kG$)
- $\beta \geq 1$ has one real part and $(\beta-1)$ “Ghost” parts

Introductory Theory

- There is an advanced theory emerging
 - A β -ghost is a spherically symmetric random variable defined on R^β
 - A shadow is a derived real or complex quantity

Wishart Matrices (arbitrary covariance)

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- For $\beta=1,2,4$, the joint eigenvalue density of A has a formula:

$$c_{m,n,\Sigma,\beta} \prod_{i=1}^n \lambda_i^{\frac{m-n+1}{2}\beta-1} \prod_{j < k} |\lambda_k - \lambda_j|^\beta \cdot {}_0F_0^{(\beta)}\left(-\frac{\beta}{2}\Lambda, \Sigma^{-1}\right),$$

where $\Lambda = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_n)$.

Joint Eigenvalue density of $G'G \Sigma$

$$c_{m,n,\Sigma,\beta} \prod_{i=1}^n \lambda_i^{\frac{m-n+1}{2}\beta-1} \prod_{j<k} |\lambda_k - \lambda_j|^\beta \cdot {}_0F_0^{(\beta)}\left(-\frac{\beta}{2}\Lambda, \Sigma^{-1}\right),$$

where $\Lambda = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_n)$.

The “ ${}_0F_0$ ” function is a hypergeometric function of two matrix arguments that depends only on the eigenvalues of the matrices. Formulas and software exist.

Generalization of Laguerre

- Laguerre:

$$C_{m,n,\beta} \prod_{i=1}^n \lambda_i^{\frac{m-n+1}{2}\beta-1} \prod_{j<k} |\lambda_k - \lambda_j|^\beta \cdot \exp\left(-\frac{\beta}{2} \sum_{i=1}^n \lambda_i\right)$$

- Versus Wishart:

$$C_{m,n,\Sigma,\beta} \prod_{i=1}^n \lambda_i^{\frac{m-n+1}{2}\beta-1} \prod_{j<k} |\lambda_k - \lambda_j|^\beta {}_0F_0^{(\beta)}\left(-\frac{\beta}{2}\Lambda, \Sigma^{-1}\right)$$

General β ?

The joint density:

$$c_{m,n,\Sigma,\beta} \prod_{i=1}^n \lambda_i^{\frac{m-n+1}{2}\beta-1} \prod_{j<k} |\lambda_k - \lambda_j|^\beta \cdot {}_0F_0^{(\beta)}\left(-\frac{\beta}{2}\Lambda, \Sigma^{-1}\right),$$

is a probability density for all $\beta > 0$.

Goals:

- Algorithm for sampling from this density
- Get a feel for the density's “ghost” meaning

Main Result

- An algorithm derived from ghosts that samples eigenvalues
- A MATLAB implementation that is consistent with other beta-ized formulas
 - Largest Eigenvalue
 - Smallest Eigenvalue

Real quantity

Working with Ghosts

$$\begin{bmatrix} g_{1,1} & g_{1,2} & g_{1,3} \\ g_{2,1} & g_{2,2} & g_{2,3} \\ g_{3,1} & g_{3,2} & g_{3,3} \end{bmatrix} \rightarrow \begin{bmatrix} \chi_{3\beta} & g_{1,2} & g_{1,3} \\ 0 & g_{2,2} & g_{2,3} \\ 0 & g_{3,2} & g_{3,3} \end{bmatrix} \rightarrow$$

$$\begin{bmatrix} \chi_{3\beta} & \chi_{2\beta} & 0 \\ 0 & g_{2,2} & g_{2,3} \\ 0 & g_{3,2} & g_{3,3} \end{bmatrix} \rightarrow \begin{bmatrix} \chi_{3\beta} & \chi_{2\beta} & 0 \\ 0 & \chi_{2\beta} & g_{2,3} \\ 0 & 0 & g_{3,3} \end{bmatrix} \rightarrow$$

$$\begin{bmatrix} \chi_{3\beta} & \chi_{2\beta} & 0 \\ 0 & \chi_{2\beta} & \chi_{\beta} \\ 0 & 0 & g_{3,3} \end{bmatrix} \rightarrow \begin{bmatrix} \chi_{3\beta} & \chi_{2\beta} & 0 \\ 0 & \chi_{2\beta} & \chi_{\beta} \\ 0 & 0 & \chi_{\beta} \end{bmatrix}$$

More practice with Ghosts

$$\begin{bmatrix} \sigma_1 g_{1,1} & \sigma_2 g_{1,2} \\ \sigma_1 g_{2,1} & \sigma_2 g_{2,2} \\ \sigma_1 g_{3,1} & \sigma_2 g_{3,2} \end{bmatrix} \longrightarrow \begin{bmatrix} \sigma_1 \chi_{3\beta} & \sigma_2 g_{1,2} \\ 0 & \sigma_2 g_{2,2} \\ 0 & \sigma_2 g_{3,2} \end{bmatrix},$$

The Algorithm for $Z=G\Sigma^{1/2}$

$$Z \equiv \begin{bmatrix} g_{1,1} & \cdots & g_{1,n} \\ \vdots & & \vdots \\ g_{m,1} & \cdots & g_{m,n} \end{bmatrix} \cdot \begin{bmatrix} d_1 & & \\ & \ddots & \\ & & d_n \end{bmatrix}$$

The Algorithm for $Z=G\Sigma^{1/2}$

$$Z \equiv \begin{bmatrix} g_{1,1} & \cdots & g_{1,n} \\ \vdots & & \vdots \\ g_{m,1} & \cdots & g_{m,n} \end{bmatrix} \cdot \begin{bmatrix} d_1 & & \\ & \ddots & \\ & & d_n \end{bmatrix}$$

By recursion, we can assume the first $n - 1$ columns of Z have a Ghost SVD:

$$Z = \begin{bmatrix} & d_n g_{1,n} \\ UTV^t & \vdots \\ & d_n g_{m,n} \end{bmatrix}$$

The real diagonal matrix T is assumed ordered.

Removing U and V

$$Z = \begin{bmatrix} & d_n g_{1,n} \\ UTV^t & \vdots \\ & d_n g_{m,n} \end{bmatrix}$$

We can use the assumptions to get rid of U and V without affecting the Ghost Spectrum of Z :

$$Z = \begin{bmatrix} t_1 & & d_n g_{1,n} \\ & \ddots & \vdots \\ & & t_{n-1} & d_n g_{n-1,n} \\ & & & \vdots \\ & & & d_n g_{m,n} \end{bmatrix}$$

Algorithm cont.

We can multiply on the left by a Ghost Orthogonal matrix of "Ghost Signs," $\text{diag}(\phi_1, \dots, \phi_m)$, to create real numbers at the expense of making the t_i 's Ghost:

$$Z = \begin{bmatrix} t_1\phi_1 & & & d_n\chi_\beta \\ & \ddots & & \vdots \\ & & t_{n-1}\phi_{n-1} & d_n\chi_\beta \\ & & & \vdots \\ & & & d_n\chi_\beta \end{bmatrix}$$

Completion of Recursion

We can multiply on the right by more Ghost Signs to make the whole matrix real:

$$Z = \begin{bmatrix} t_1 & & & d_m \chi_\beta \\ & \ddots & & \vdots \\ & & t_{m-1} & d_m \chi_\beta \\ & & & \vdots \\ & & & d_m \chi_\beta \end{bmatrix}$$

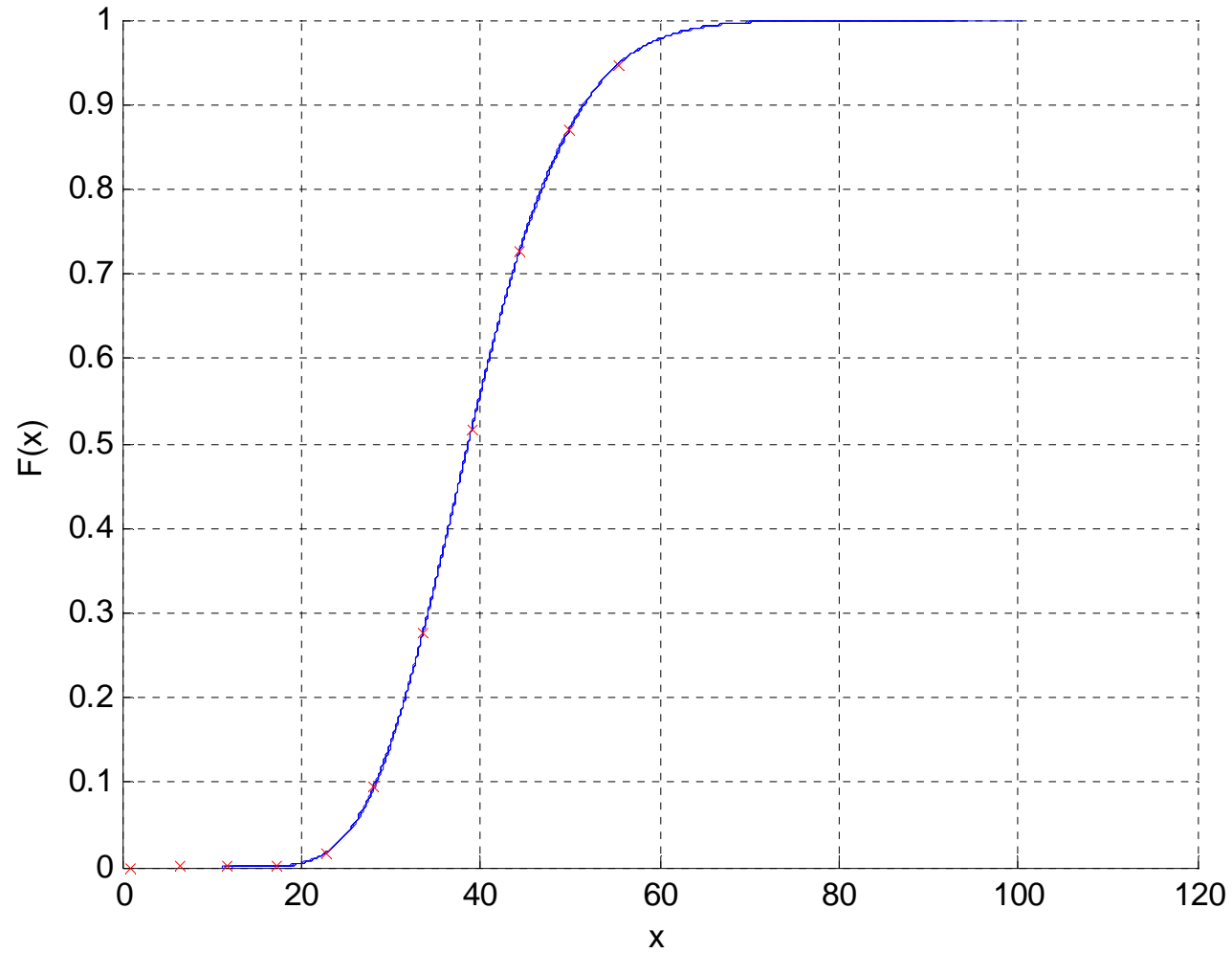
We can take the real SVD of this matrix, completing the recursion.

Numerical Experiments – Largest Eigenvalue

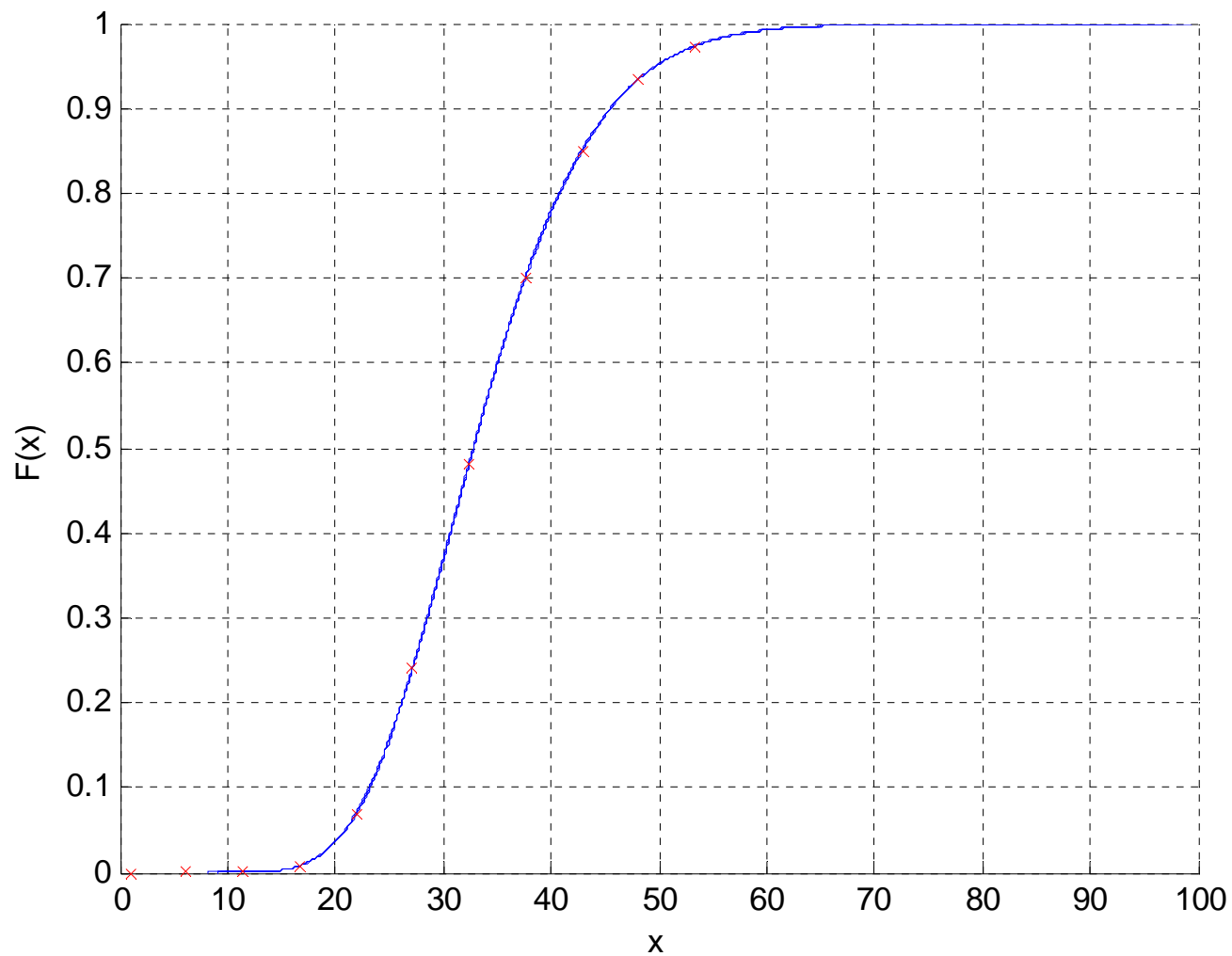
- Analytic Formula for largest eigenvalue dist

$$c_{m,n,\beta} \cdot \det\left(\frac{x}{2}\beta\Sigma^{-1}\right)^{\frac{m}{2}\beta} \cdot {}_1F_1^{(\beta)}\left(\frac{m}{2}\beta; \frac{m+n-1}{2}\beta+1; -\frac{x}{2}\beta\Sigma^{-1}\right).$$

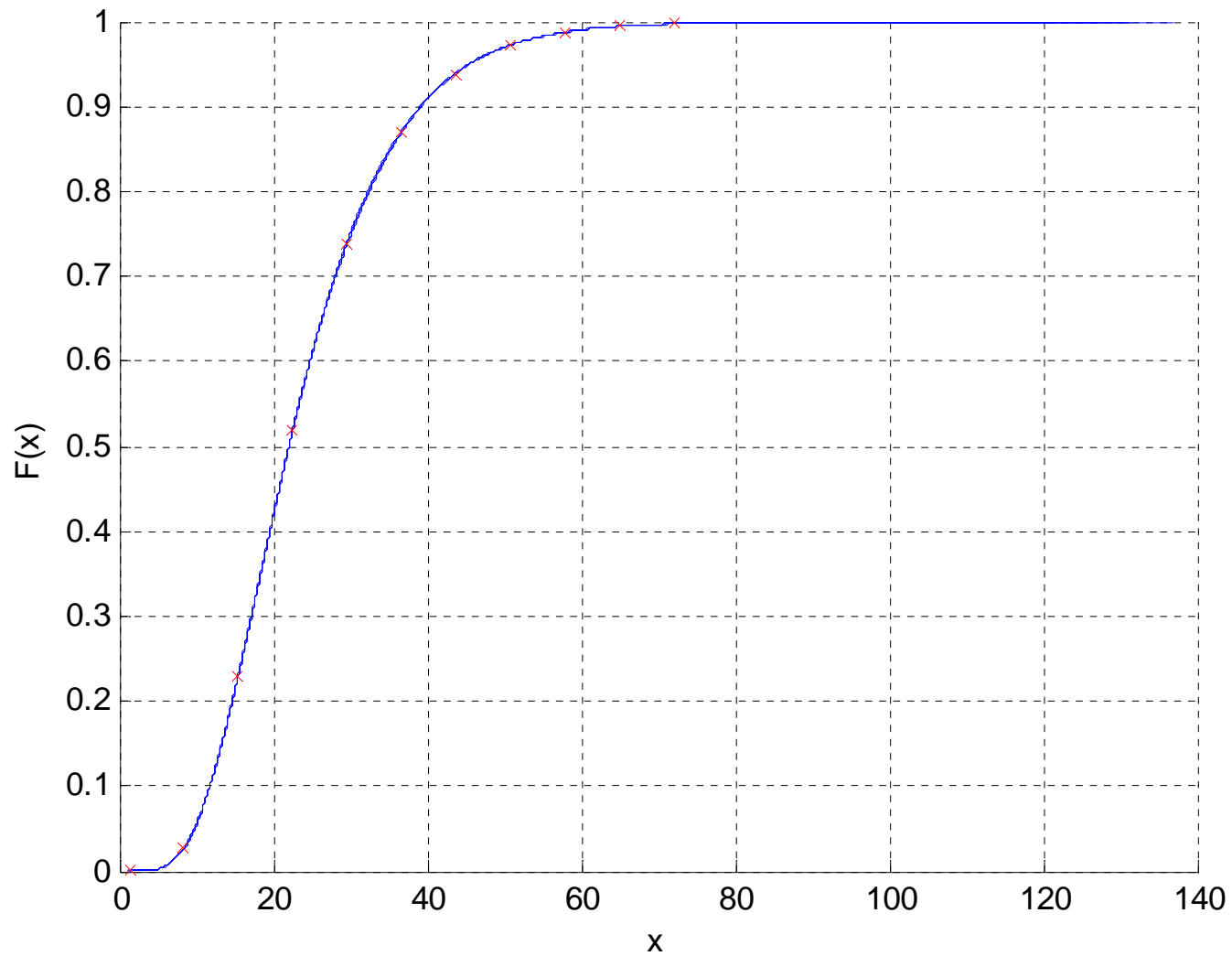
- E and Koev: software to compute



$$m = 3, n = 3, \beta = 5, \Sigma = \text{diag}(1.1, 1.2, 1.4)$$



$$m = 4, n = 4, \beta = 2.5, \Sigma = \text{diag}(1.1, 1.2, 1.4, 1.8)$$



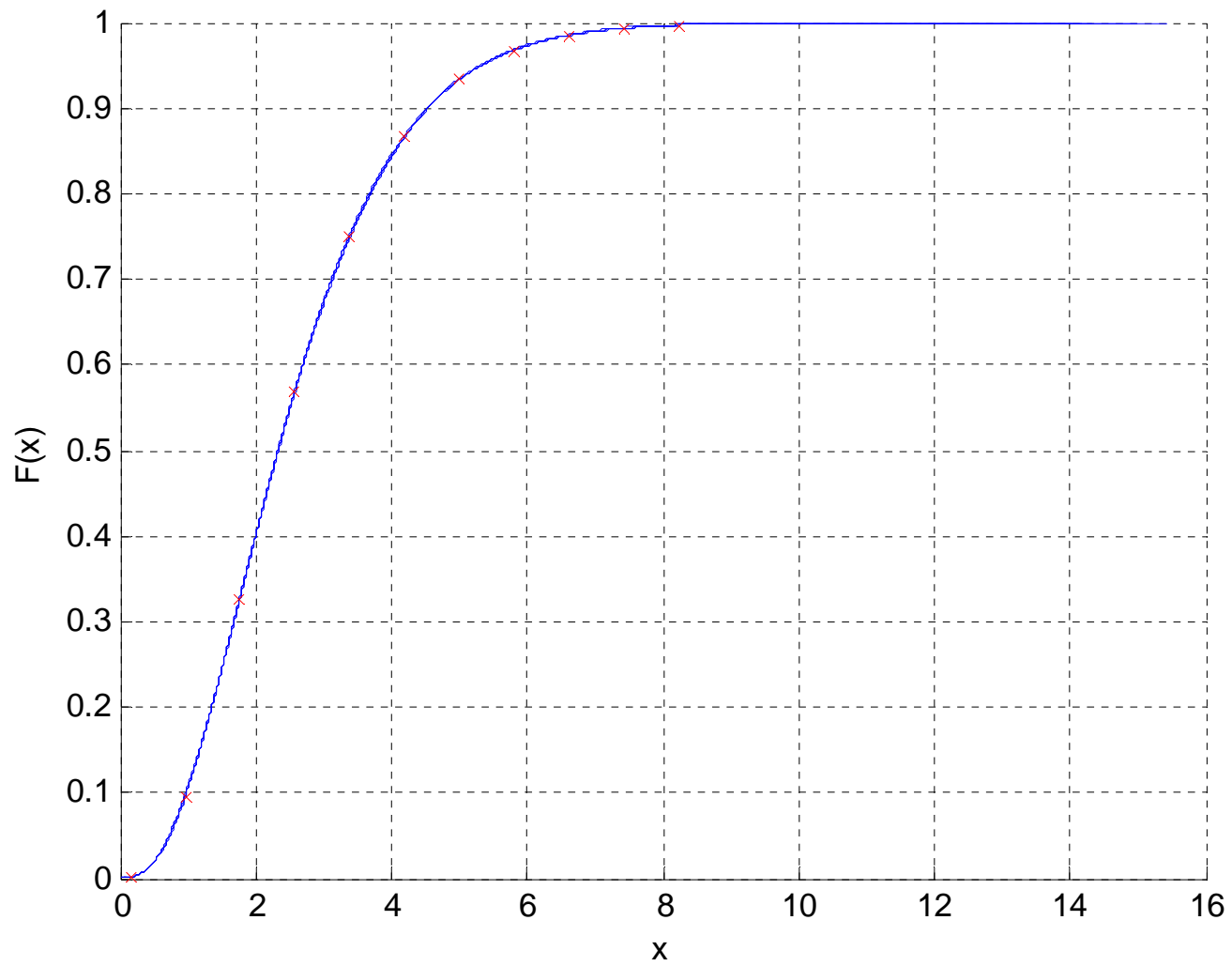
$$m = 5, n = 4, \beta = 0.75, \Sigma = \text{diag}(1, 2, 3, 4) \quad 37/47$$

Smallest Eigenvalue as Well

The cdf of the smallest eigenvalue,

for integer values of $r = (n - m + 1)\beta/2 - 1$, is

$$P(x) = 1 - \exp\left(\operatorname{tr}\left(\frac{x}{2}\Sigma^{-1}\right)\right) \sum_{k=0}^{mp} \widehat{\sum}_{\kappa} \frac{C_{\kappa}^{(\beta)}\left(\frac{x}{2}\Sigma^{-1}\right)}{k!}.$$



$$m = 5, n = 4, \beta = 3, \Sigma = \text{diag}(1.1, 1.2, 1.4, 1.8)$$

Goals

- Continuum of Haar Measures generalizing orthogonal, unitary, symplectic
- New Definition of Jack Polynomials generalizing the zonals
- Computations! E.g. Moments of the Haar Measures
- Place finite random matrix theory “ β ” into same framework as infinite random matrix theory: specifically β as a knob to turn down the randomness, e.g. Airy Kernel
– $d^2/dx^2+x+(2/\beta^{1/2})dW \leftarrow$ White Noise

Formally

- Let $S_n = 2\pi/\Gamma(n/2)$ = “surface area of sphere”
 - Defined at any $n = \beta > 0$.
- A β -ghost x is formally defined by a function $f_x(r)$ such that $\int_{r=0}^{\infty} f_x(r) r^{\beta-1} S_{\beta-1} dr = 1$.
- Note: For β integer, the x can be realized as a random spherically symmetric variable in β dimensions
- Example: A β -normal ghost is defined by
 - $f(r) = (2\pi)^{-\beta/2} e^{-r^2/2}$
- Example: Zero is defined with constant $\delta(r)$.
- Can we do algebra? Can we do linear algebra?
- Can we add? Can we multiply?

A few more operations

- $\|x\|$ is a real random variable whose density is given by $f_x(r)$
- $(x+x')/2$ is real random variable given by multiplying $\|x\|$ by a beta distributed random variable representing a coordinate on the sphere

Addition of Independent Ghosts:

- Addition returns a spherically symmetric object
- Have an integral formula
- Prefer: Add the real part, imaginary part completed to keep spherical symmetry

Multiplication of Independent Ghosts

- Just multiply $\|z\|^2$'s and plug in spherical symmetry
- Multiplication is commutative
 - (Important Example: Quaternions don't commute, but spherically symmetric random variables do!)

Understanding $\prod |\lambda_i - \lambda_j|^\beta$

- Define volume element $(dx)^\wedge$ by
 $(r dx)^\wedge = r^\beta (dx)^\wedge$ (β -dim volume, like fractals, but don't really see any fractal theory here)
- Jacobians: $A = Q \Lambda Q'$ (Sym Eigendecomposition)
 $Q' dA Q = d\Lambda + (Q' dQ) \Lambda - \Lambda (Q' dQ)$
 $(dA)^\wedge = (Q' dA Q)^\wedge = \text{diagonal}^\wedge \wedge \text{strictly-upper diagonal}$
 $\text{diagonal} = \prod d\lambda_i = (d\Lambda)^\wedge$
 $\text{off-diag} = \prod \left((Q' dQ)_{ij} (\lambda_i - \lambda_j) \right)^\wedge = (Q' dQ)^\wedge \prod |\lambda_i - \lambda_j|^\beta$

Haar Measure

$$E(\text{trace}(AQBQ')^k) = \sum C_k(A)C_k(B)/C_k(I)$$

Example: Suppose you know C_k 's a-priori.

minimants (Think

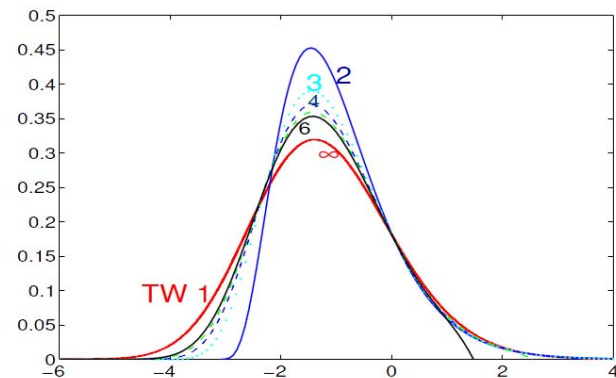
$E(q_{11} ^4)$	$=$	$\frac{1+\alpha}{(n+\alpha)n}$	$E(q_{11} ^4 q_{12} ^2)$	$=$	$\frac{1+\alpha}{n(n+\alpha)(n+2\alpha)}$
$E(q_{11} ^2 q_{12} ^2)$	$=$	$\frac{1}{(n+\alpha)n}$	$E(q_{11} ^2 q_{12} ^2 q_{13} ^2)$	$=$	$\frac{1}{n(n+\alpha)(n+2\alpha)}$
$E(q_{11} ^2 q_{22} ^2)$	$=$	$\frac{n+\alpha-1}{n+\alpha}$	$E(q_{11} ^2 q_{12} ^2 q_{23} ^2)$	$=$	$\frac{(n-1)n(n+\alpha)(n+2\alpha)}{(1+\alpha)(n+2\alpha-1)}$
$E(q_{11} ^6)$	$=$	$\frac{n(n-1)(n+\alpha)}{(1+\alpha)(1+2\alpha)}$	$E(q_{11} ^4 q_{22} ^2)$	$=$	$\frac{(n-1)n(n+\alpha)(n+2\alpha)}{(1+\alpha)(n+2\alpha-1)}$

- Then can...
- Example: $E(|q_{11}|^2|q_{22}|^2)$
- $\alpha := 2/\beta$
- Can Gram-Schmidt the ghosts. Same as...
- coming up!

Further Uses of Ghosts

- Multivariate Orthogonal Polynomials
- Tracy-Widom Laws
- Largest Eigenvalues/Smallest Eigenvalues

$$A_p \sim W_p(n, I); \quad n/p = 5; \quad (\lambda_{\max}(A_p) - \mu_p)/\sigma_p \rightarrow \text{TW}_1$$



- Expect lots of uses to be discovered...